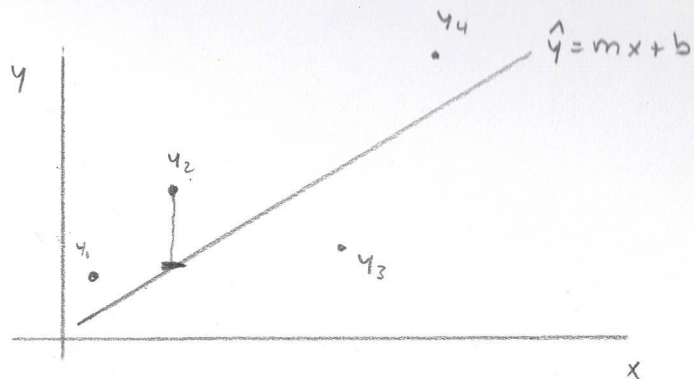


Linear Regression



n observations of y

y : income

x : years of college

(explanatory variable)

Q: (1) degree of correlation?

(2) best (linear) fit?

SSE: sum squared error

$$\sum (\hat{y}_i - y_i)^2 = \|\vec{\hat{y}} - \vec{y}\|^2 \leftarrow \text{norm}$$

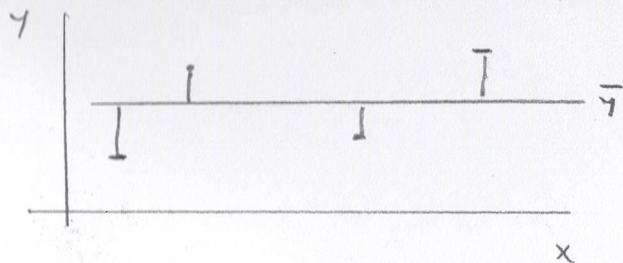
$$\|\vec{x}\| = \sqrt{x_1^2 + x_2^2 + \dots + x_n^2}$$

$SSE = 0 \Rightarrow \hat{y}_i = y_i$ for all $i \Rightarrow$ perfect fit

SST: sum squared total "error if using mean as predictor"

$$\sum (y_i - \bar{y})^2$$

$\nwarrow$ mean of all observations



$SST = 0 \Rightarrow$ all values equal mean

$$R^2 = 1 - \frac{SSE}{SST} \quad \text{correlation coefficient}$$

$SSE = 0 \Rightarrow$ perfect fit $R^2 = 1$

$SSE = SST \Rightarrow$ no correlation $R^2 = 0$

"no linear model is better than mean"

What is the best linear model?

$$\hat{y} = mx + b$$

predictor model

$$\left. \begin{aligned} \hat{y}_1 &= mx_1 + b \\ \hat{y}_2 &= mx_2 + b \\ &\vdots \\ \hat{y}_n &= mx_n + b \end{aligned} \right\} \begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \vdots \\ \hat{y}_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \begin{bmatrix} b \\ m \end{bmatrix}$$

$$\hat{\mathbf{y}} = \mathbf{A} \boldsymbol{\beta}$$

"best fit" $\Rightarrow$ minimize SSE

$$\text{minimize } \|\hat{\mathbf{y}} - \mathbf{y}\|^2$$

$$\text{minimize } \|\mathbf{A}\boldsymbol{\beta} - \mathbf{y}\|^2 \Rightarrow \text{find } \boldsymbol{\beta} \text{ that minimizes SSE}$$

steps $\frac{\partial \text{SSE}}{\partial \boldsymbol{\beta}} = 0 \Rightarrow \text{solve for } \boldsymbol{\beta}$

$$SSE = \|A\beta - y\|^2$$

$$= (A\beta - y)^T (A\beta - y)$$

$$= (\beta^T A^T - y^T) (A\beta - y)$$

$$= \beta^T A^T A \beta - y^T A \beta - \underbrace{\beta^T A^T y + y^T y}_{\substack{(A\beta)^T y \\ y^T A \beta} \text{ transpose}}$$

$$= \beta^T A^T A \beta - 2y^T A \beta + y^T y$$

$$\frac{\partial SSE}{\partial \beta} = 2$$

derivative rules

$$\frac{\partial \vec{y}}{\partial \vec{x}} \stackrel{\text{def}}{=} \begin{bmatrix} \frac{\partial y_1}{\partial x_1} & \dots & \frac{\partial y_n}{\partial x_n} \\ \vdots & & \vdots \\ \frac{\partial y_1}{\partial x_m} & \dots & \frac{\partial y_n}{\partial x_m} \end{bmatrix}$$

useful identities

$$\|x^2\| = \sum x_i^2 = x^T x$$

$$\|x\| = \sqrt{\sum x_i^2}$$

$$\begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix} [x_1 \dots x_n] = x_1^2 + \dots + x_n^2$$

$$(A^T)^T = A$$

$$(AB)^T = B^T A$$

$$(A+B)^T = A^T + B^T$$

$$SSE = \beta^T A^T A \beta - 2 y^T A \beta + y^T y$$

$$\frac{\partial SSE}{\partial \beta} = 2 A^T A \beta - 2 A^T y + 0 = 0$$

$$A^T A \beta = A^T y$$

$$(A^T A)^{-1} (A^T A) \beta = (A^T A)^{-1} A^T y$$

$$\boxed{\beta = (A^T A)^{-1} A^T y}$$

note: $A^T A$ is always invertible if column vectors are linearly independent

Common matrix derivatives

$\vec{y}$	$\frac{\partial \vec{y}}{\partial \vec{x}}$
$A \vec{x}$	A
$\vec{x}^T A$	A^T
$\vec{x}^T \vec{x}$	$2\vec{x}$
$\vec{x}^T A^T A \vec{x}$	$2 A^T A \vec{x}$
$\vec{y}^T A \vec{x}$	$A^T \vec{y}$

Linear regression in practice

X	Y
1	-0.2
2	2.2
3	3.5
⋮	⋮
8	12

column of 1

column of x_i

vector of y

$$A = \begin{bmatrix} 1 & -0.2 \\ 1 & 2.2 \\ \vdots & \vdots \\ 1 & 8 \end{bmatrix}$$

$$y = \begin{bmatrix} -0.2 \\ 2.2 \\ \vdots \\ 12 \end{bmatrix}$$

n explanatory
variable

regressed
variable

$$\beta = \text{inv}(A^T A) A y = \begin{bmatrix} m \\ b \end{bmatrix}$$



defined
in any language

may need some

Linear Algebra package

$$SSE = \text{norm}(Ax - y)^2$$

$$SST = \text{norm}(y - \text{mean}(y))^2$$

$$R^2 = 1 - \frac{SSE}{SST}$$

may need Statistics package

Multilinear Regression

base case $y = \frac{1}{\text{mpg}}$

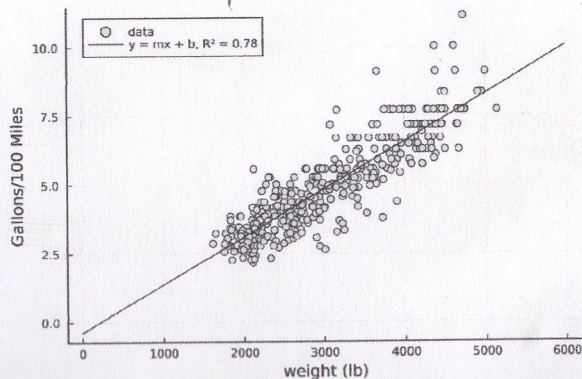
(Gallons per 100 miles)

$x = \text{weight}$

$$y = mx + b$$

$$R^2 = 0.79$$

	mpg	cylinders	displacement	hp	weight	acceleration	year	mc
1	18.0	8	307.0	130.0	3504.0	17.0	70	"chevrolet ch
2	15.0	8	350.0	165.0	3693.0	11.5	70	"buick skylari
3	18.0	8	318.0	150.0	3436.0	11.0	70	"plymouth sat
4	16.0	8	304.0	160.0	3433.0	12.0	70	"amc rebel ss
5	17.0	8	302.0	140.0	3449.0	10.5	70	"ford torino"
6	15.0	8	429.0	198.0	4341.0	10.0	70	"ford galaxie
7	14.0	8	454.0	220.0	4364.0	9.0	70	"chevrolet im
8	14.0	8	440.0	215.0	4312.0	8.5	70	"plymouth fur
9	14.0	8	455.0	225.0	4425.0	10.0	70	"pontiac cata
10	15.0	8	390.0	190.0	3850.0	8.5	70	"amc ambassad
: more								
392	31.0	4	119.0	82.0	2720.0	19.4	82	"chevy s-10"



"cars" Dataset

Multivariable model

$$y = c_1 x_1 + c_2 x_2 + c_3 x_3 + c_4 \text{ "intercept"}$$

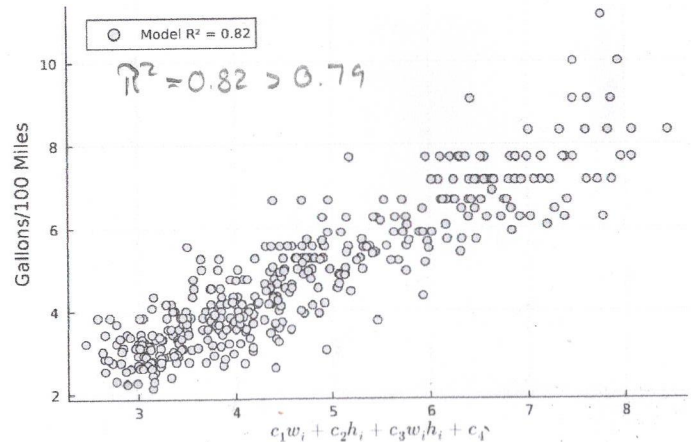
$\uparrow$ weight $\uparrow$ horse power $\nwarrow$ weight $\times$ horse power

$$\hat{y} = A \beta$$

$$\begin{bmatrix} \hat{y}_1 \\ \vdots \\ \hat{y}_n \end{bmatrix} = \begin{bmatrix} x_{11} & x_{21} & x_{31} & 1 \\ \vdots & \vdots & \vdots & \vdots \\ x_{1n} & x_{2n} & x_{3n} & 1 \end{bmatrix} \begin{bmatrix} c_1 \\ c_2 \\ c_3 \\ c_4 \end{bmatrix}$$

Adding variables increases
explanatory power

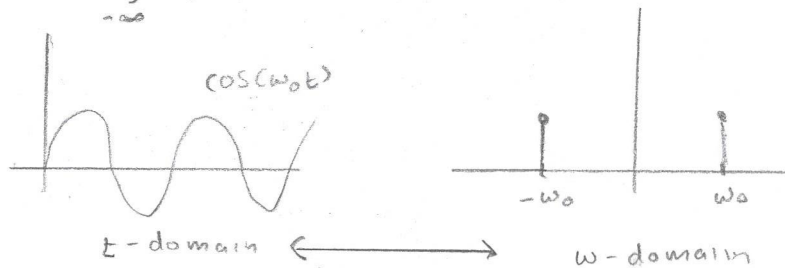
$$\beta = (A^T A)^{-1} A^T y$$



Integral Equations (Domain Transformations)

Fourier Transform

$$F(\omega) = \int_{-\infty}^{\infty} \exp(-i\omega t) f(t) dt$$



Laplace Transform

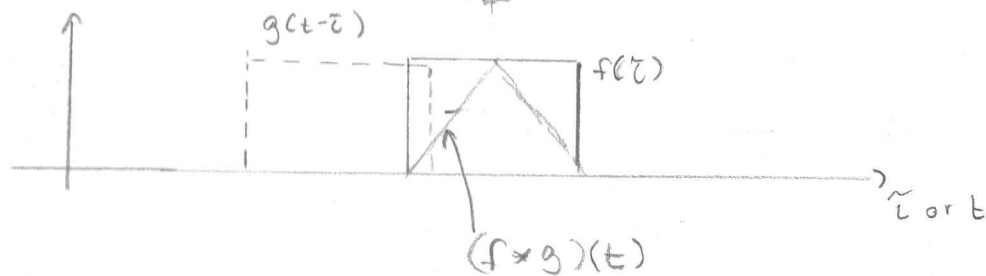
$$F(s) = \int_{-\infty}^{\infty} \exp(-st) f(t) dt$$

s is $a+bi$ (complex) number

t-domain $\longleftrightarrow$ s-domain

Convolution

$$(f * g)(t) = \int_{-\infty}^{\infty} g(t-\tau) f(\tau) d\tau$$



General Integral Equation

$$y(x) = \int_a^x k(x,t) f(t) dt$$

limits
can differ from
 $\pm \infty$

↑
kernel function

$$k(x,t) = g(x-t) \Rightarrow \text{convolution}$$

$$k(x,t) = \exp(-xt) \Rightarrow \text{Laplace}$$

$$k(x,t) = \exp(-ixt) \Rightarrow \text{Fourier}$$

⋮

t - domain



x - domain

Applications / Problem

- Measurement is in the x -domain $\Rightarrow$ quantity of interest is in t -domain

- surface temperature evolution
- size distribution function
- Optical measurements / remote sensing
- FTIR

⋮

- Inversion $\Rightarrow$ find $f(t)$ from discrete $y(x)$

Example Inversion Problem

$$y(x) = \int_a^b k(x,t) u(t) dt$$

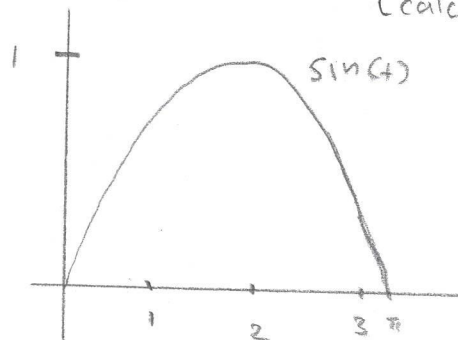
"Fredholm Integral Eq."

$$k(x,t) = \exp(x \cdot \cos(t)) \quad u(t) = \sin(t) \quad y(x) = \frac{2 \sinh(x)}{x}$$

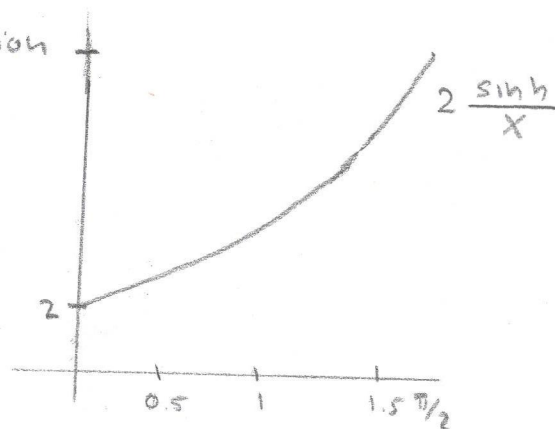
$$a=0, b=\pi \quad t \in [0, \pi] \quad \text{and} \quad x \in [0, \pi/2]$$

test : $\int_0^{\pi} \exp(x \cos(t)) \sin(t) dt = \frac{2 \sinh x}{x}$

→ verify using integration (calculus)



t -domain



x -domain

Discretized Integral Equation

$$\int_a^b k(x, t) u(t) dt = y(x)$$

$$A \vec{u} = \vec{y}$$

$$\vec{u} = [u_1, \dots, u_n]$$

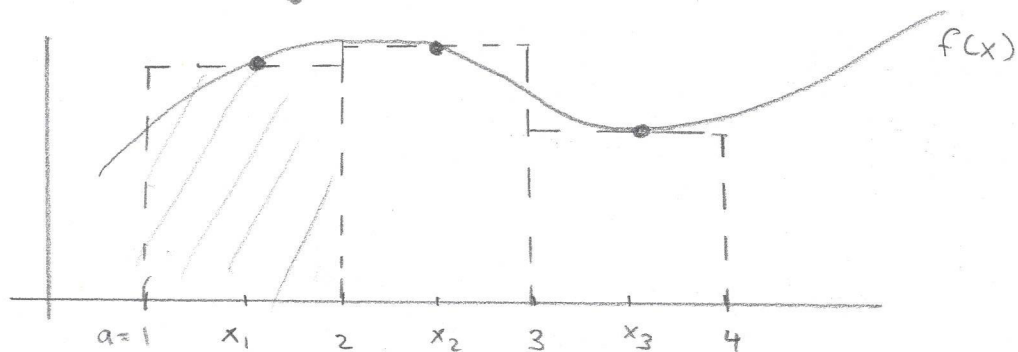
$$\vec{y} = [y_1, \dots, y_m]$$

$$A = \mathbb{R}^{m \times n}$$

Question: how to find A ? (next page)

$$\sum_{i=1}^n w_n k(x_j, t_i) u(t_i) \approx y_j$$

Numerical Integration



$n=3$: number of data points

$a=1$, $b=4$: limits of integration

$w = \frac{b-a}{n}$ bin width (here $w=1$)

$$\int_a^b f(x) dx = \sum_{i=1}^n f(x_i) w$$

} quadrature algorithm

$$x_i = (i - \frac{1}{2})w + a$$

$$x_1 = (1 - 0.5)1 + 1 = 1.5$$

$$x_2 = (2 - 0.5)1 + 1 = 2.5$$

$$x_3 = (3 - 0.5)1 + 1 = 3.5$$

$$\int_a^b k(x, t) u(t) dt = y(x) \quad t \in [a, b] \quad x \in [c, d]$$

u is discretized into n points $i = 1 \dots n$ $w_n = \frac{b-a}{n}$

y is discretized into m points $j = 1 \dots m$ $w_m = \frac{d-c}{m}$

$$y_i = \int_a^b k(x_i, t) u(t) dt$$

$$y_i = k(x_i, t_1) u(t_1) w_n + k(x_i, t_2) u(t_2) w_n + \dots + k(x_i, t_n) u(t_n) w_n$$

$$y_m = k(x_m, t_1) u(t_1) w_n + k(x_m, t_2) u(t_2) w_n + \dots + k(x_m, t_n) u(t_n) w_n$$

$$\begin{bmatrix} y_1 \\ \vdots \\ y_m \end{bmatrix} = \begin{bmatrix} k(x_1, t_1) w_n & \dots & k(x_1, t_n) w_n \\ \vdots & & \vdots \\ k(x_m, t_1) w_n & \dots & k(x_m, t_n) w_n \end{bmatrix} \begin{bmatrix} u_1 \\ \vdots \\ u_n \end{bmatrix}$$

$$a_{j,i} = w_n k(x_j, t_i)$$

$$A = \mathbb{R}^{m \times n}$$

$$y = A u$$

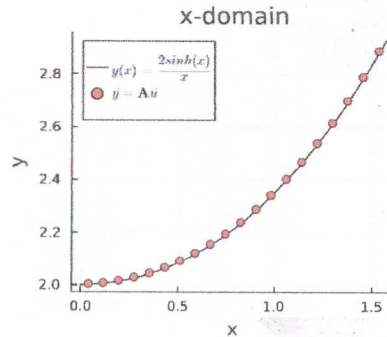
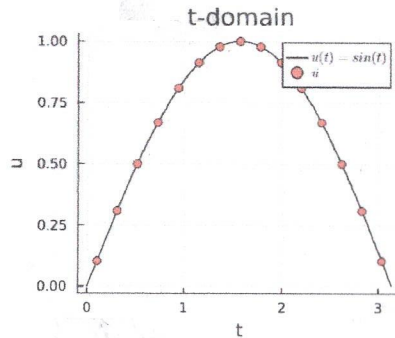
$\mathbb{R}^m$

$u \in \mathbb{R}^n$

Example Discretization

→ n, m can be independently varied

→ overdetermined and underdetermined transforms

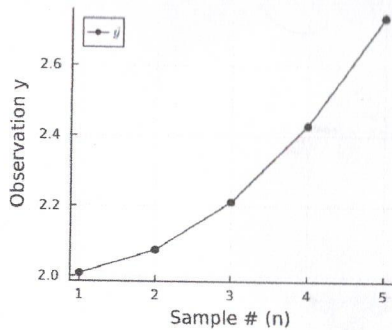


$n = 15$

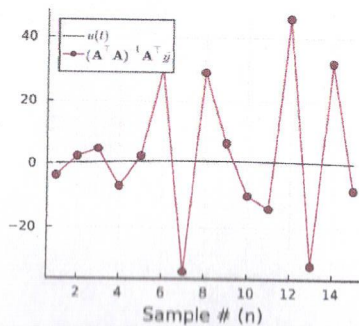
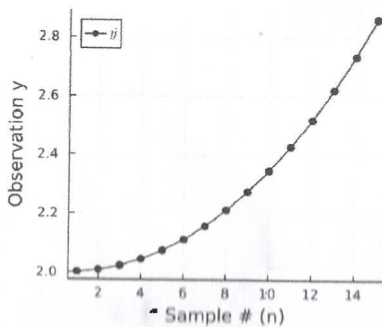
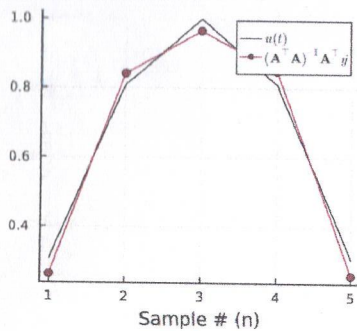
$m = 20$

Example Inversion

Observation in
X-domain



Prediction in
t-domain



$$y = A u(t)$$

SSE: minimize $\|Au - y\|^2$
 $\uparrow$
 find $\vec{u}$
 that is best

$$u = (A^T A)^{-1} A^T y$$

works well for small n

fails for large n
 why?

III posed Problems

- Lack of a unique solution
- Sensitivity to initial conditions (noise) $\rightarrow$ weather prediction
- nonexistence of a solution

$$\begin{array}{l} A = \begin{bmatrix} 1 & 2 \\ 0 & 0 \end{bmatrix} \quad \det A = 0 \\ A = \begin{bmatrix} 1 & 2 \\ 2 & 4 \end{bmatrix} \quad \det A = 0 \end{array} \left\{ \begin{array}{l} \text{- underdetermined} \\ \text{- no unique solution} \\ A^{-1} \text{ not defined} \end{array} \right.$$

$\text{rank}(A) =$ maximum number of linearly independent columns

here $\text{rank}(A) = 1$

Full rank = 2 (max possible)

A is "rank deficient"

Moore - Penrose Inverse

$$A^+ = (A^T A)^{-1} A^T$$

- (1) for $m=n$ $A^+ = A^{-1}$ if A is full rank
 - (2) $m > n$:
 - overdetermined system (e.g. linear regression case)
 - more eq. than unknowns
 - minimizes $\|A u - y\|^2$
 - (3) $n > m$:
 - underdetermined system
 - fewer eq. than unknowns
 - infinite solutions
- if A is not full rank A^+ must be computed using SVD
 - use `pinv(A)` in most languages

Singular Value Decomposition

Full SVD

singular values

$$A = U \Sigma V^T$$

$\mathbb{R}^{m \times n}$ $\mathbb{R}^{m \times m}$ $\mathbb{R}^{n \times n}$

left singular vectors right singular vectors

$$\Sigma = \begin{bmatrix} \sigma_1 & 0 & \dots & \dots & \dots & 0 \\ 0 & \ddots & & & & \\ \vdots & & \ddots & & & \\ 0 & & & \ddots & & \\ & & & & \sigma_p & 0 \\ & & & & & \ddots \\ & & & & & & 0 \end{bmatrix}$$

$$\sigma_1 > \sigma_2 > \dots > \sigma_p$$

$$p = \min \{m, n\}$$

- can find using numerical methods
- $O(n^2) \Rightarrow$ expensive
- use $\text{svd}(A)$ from linear algebra

Moore - Penrose Inverse

$$A^+ = V \Sigma^+ U^T$$

$$\Sigma^+ = \begin{bmatrix} \frac{1}{\sigma_1} & 0 & - & - & - & 0 \\ 0 & \ddots & & & & \vdots \\ 0 & 0 & - & - & - & \frac{1}{\sigma_p} \\ 0 & 0 & - & - & - & 0 \end{bmatrix}$$

- numerical method to find A^+
- if $\sigma_i = 0$ then a zero is placed
- if A is rank deficient one or more $\sigma_i = 0$

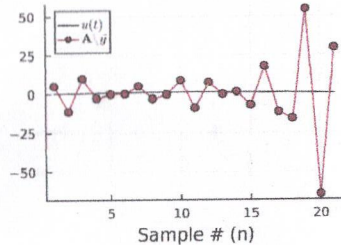
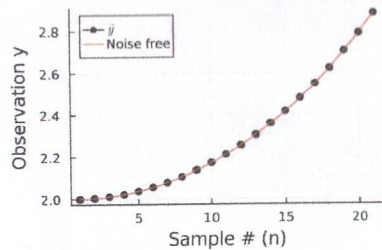
Inversion Revisited

$$\frac{1}{10^{-15}} = 10^{15} \Rightarrow \text{large entries in } \Sigma^+$$

"infected by numerical or measurement noise"

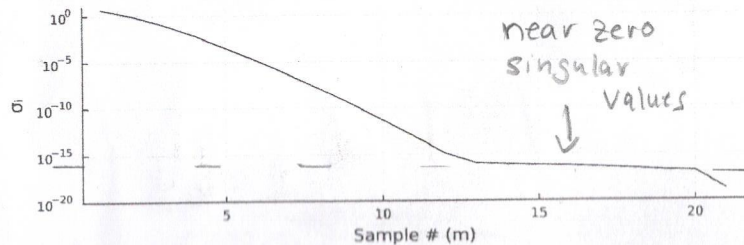
x-domain

t-domain



$$u = (A^T A)^{-1} A^T y$$

$$u = A^+ y$$



log scale

machine precision
for 64 bit floating point
 $\approx 10^{-16}$

no new information

Regularized Inverse (Noise Filtering)

$$u_R = \text{minimize} \left\{ \underbrace{\|A u - y\|^2}_{\text{ordinary regression}} + \underbrace{\lambda^2 \|L(u - u_0)\|^2}_{\substack{\uparrow \\ \text{regularization parameter}}} \right\}$$

filter matrix
initial guess

$$u_R = (A^T A + \lambda^2 L^T L)^{-1} (A^T y + \lambda^2 L^T L u_0)$$

$$u_0 = A^+ y \quad !$$

$$\lim_{\lambda \rightarrow \infty} u_R = u_0$$

Starting point for $L = \underline{I}$

$\uparrow$
identity
matrix

$u_0 = 0$
(no initial guess)

then

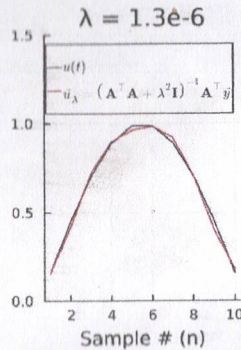
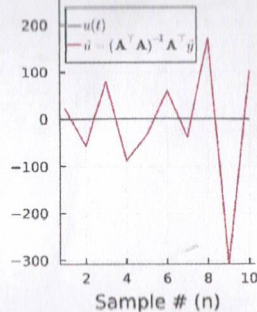
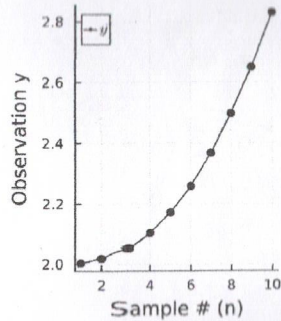
$$u_R = (A^T A + \lambda^2 I)^{-1} A^T y \quad \text{"Ridge Regression"}$$

In version Revisited

- need to try range of λ
- $\lambda = 0 \Rightarrow$ bad solution
- λ large $\Rightarrow$ approach zero $\Rightarrow$ bad solution
- a λ exists to find good inversion
- $\Rightarrow$ how to find λ ?

x-domain

t-domain



$$u_{\lambda} = (A^T A + \lambda^2 I)^{-1} A^T y$$

$\lambda = 0$

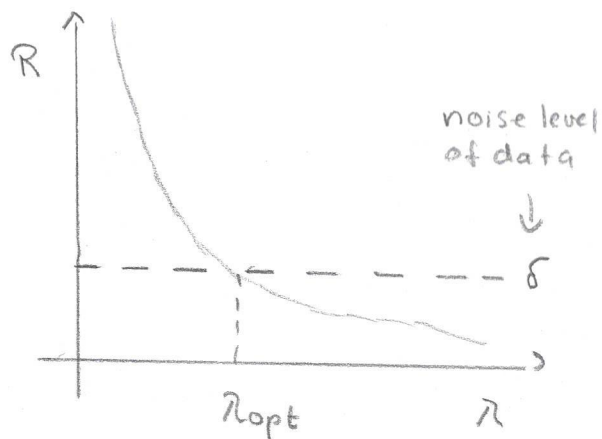
Methods to find λ

$$u_\lambda = \text{minimize} \left\{ \underbrace{\|Au - y\|^2}_{R} + \lambda^2 \underbrace{\|L(u - u_0)\|^2}_{S} \right\}$$

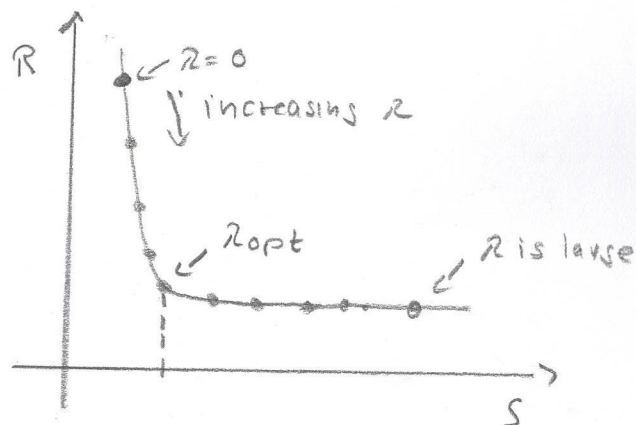
R = Residual Norm

S = Solution Norm

Morozov



L-curve



- λ_{opt} is "corner" of L-curve
- find visually or algorithm